# Financial Stability and Optimal Interest-Rate Policy

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Federal Reserve Board

#### The New Normal of Monetary Policy, March 2015

The views expressed herein are of the authors and do not represent the opinions of the Federal Reserve Board of Governors or the Federal Reserve System.

#### Motivation

Should interest-rate policy be altered in response to changes in credit conditions?

Focus on optimal policy.

- Intertemporal trade-off between costs and benefits of leaning against financial imbalances.
  - Reduced economic activity today in exchange for lower likelihood of crises tomorrow.

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How quantitatively relevant?

#### What We Do

- Build a New Keynesian model with endogenous financial crises.
- Characterize optimal interest-rate policy in presence of endogenous probability of a crisis.
  - Non-linear quadratic approach.
- Characterize optimal interest-rate policy accounting for parameter uncertainty.

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Bayesian and Robust Control approaches.

# What We Find

 Optimal interest-rate adjustment in response to credit conditions is very small under baseline calibration.

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- Optimal interest-rate adjustment in response to credit conditions is very small under baseline calibration.
- Optimal policy can call for larger interest-rate adjustments when crisis probability is more sensitive to credit imbalances or crisis is expected to be severe.
- Uncertainty over the crisis probability and its sensitivity to monetary policy calls for more aggressive interest-rate policy.
  - Deviation from attenuation principle (Brainard (1967)).

# Outline

The Model

Calibration

**Baseline Results** 

Sensitivity Analysis

Role of Uncertainty

Conclusions

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# The Model

- New Keynesian sticky-price model with an endogenous financial crisis event.
- Crisis follows a Markov process. Transition probability depends on aggregate financial conditions.
- Two periods, denoted t = 1 and t = 2.
- Trade-off: Tighter interest-rate policy in normal times can lower output in t = 1 and reduce probability of crisis occurring in period t=2.

• Output gap y, inflation  $\pi$ , and credit conditions L in t = 1:

$$y_{1} = E_{1}^{ps} y_{2} - \sigma [i_{1} - E_{1}^{ps} \pi_{2}]$$
  

$$\pi_{1} = \kappa y_{1} + E_{1}^{ps} \pi_{2}$$
  

$$L_{1} = \rho_{L} L_{0} + \phi_{i} i_{1} + \phi_{y} y_{1} + \phi_{\pi} \pi_{1} + \phi_{0}.$$

• In period t = 2 output gap and inflation can take values:

$$(y_2, \pi_2) = \begin{cases} (y_{2,nc}, \pi_{2,nc}), & \text{with probability } 1 - \gamma_1 \\ (y_{2,c}, \pi_{2,c}), & \text{with probability } \gamma_1 \end{cases}$$

with  $y_{2,c} < y_{2,nc} = 0$  and  $\pi_{2,c} < \pi_{2,nc} = 0$  and

$$\gamma_1 = \frac{\exp(h_0 + h_1 L_1)}{1 + \exp(h_0 + h_1 L_1)}$$

•  $E_1^{ps}$  non-rational private sector expectations in t = 1.

Private sector expectations are optimistic on crisis probability.
 Supporting evidence from surveys.

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 Assumption eliminates precautionary saving motive. Focus on intertemporal policy trade-off.

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# Central Bank's Problem

The policy problem: choose policy rate in t=1 given initial credit conditions, L<sub>0</sub>:

$$WL_1 = \min_{i_1} u(y_1, \pi_1) + \beta E_1[WL_2]$$

subject to the private sector equilibrium conditions.

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Per-period welfare loss:

$$u(y_1, \pi_1) = \frac{1}{2}(\lambda y_1^2 + \pi_1^2).$$

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Per-period welfare loss:

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• Expected welfare loss in period t = 2:

$$E_1[WL_2] = (1 - \gamma_1)WL_{2,nc} + \gamma_1WL_{2,c}$$

where:

$$WL_{2,nc} = u(y_{2,nc}, \pi_{2,nc}), \quad WL_{2,c} = \frac{u(y_{2,c}, \pi_{2,c})}{1 - \beta \mu}$$

# Calibration - NK block

| Parameter                               | Description                           | Parameter Value                          |  |  |  |
|---|---------------------------------------|--|--|--|--|
| Standard parameters                     |                                       |  |  |  |  |
| β                                       | Discount Factor                       | 0.995                                    |  |  |  |
| $\sigma$                                | Interest-rate sensitivity of output   | 1.0                                      |  |  |  |
| $\kappa$                                | Slope of the Phillips Curve           | 0.024                                    |  |  |  |
| $\lambda$                               | Weight on output stabilization        | 1/16                                     |  |  |  |
| i*                                      | Long-Run Natural Rate of Interest     | 0.01                                     |  |  |  |
| Parameters related to the second period |                                       |  |  |  |  |
| У2, <i>п</i> с                          | Output gap in the non-crisis state    | 0  |  |  |  |
| $\pi_{2,nc}$                            | Inflation gap in the non-crisis state | 0  |  |  |  |
| $WL_{2,nc}$                             | Loss in the non-crisis state          | 0  |  |  |  |
| <i>Y</i> 2, <i>c</i>                    | Output gap in the crisis state        | -0.1                                     |  |  |  |
| $\pi_{2,c}$                             | Inflation gap in the crisis state     | -0.02/4                                  |  |  |  |
| $\mu$                                   | Persistence of the crisis state       | 7/8                                      |  |  |  |
| $WL_{2,c}$                              | Loss in the crisis state              | $\frac{u(y_{2,c},\pi_{2,c})}{1-\beta u}$ |  |  |  |
| $\epsilon$                              | Perceived crisis probability          | 0.05/100                                 |  |  |  |

# Calibration - Crisis Probability

Calibrate transition probability parameters:

$$\gamma_1 = \frac{\exp(h_0 + h_1 L_1)}{1 + \exp(h_0 + h_1 L_1)}$$

Credit conditions:

$$L_1 = \phi_0 + \rho_L L_0 + \phi_y y_1 + \phi_i i_1 + \phi_\pi \pi_1 + \varepsilon_1$$

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 Adapt Schularick and Taylor (2012) findings: L cumulative 5-year growth rate of real bank loans.

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- Adapt Schularick and Taylor (2012) findings: L cumulative 5-year growth rate of real bank loans.
- Growth of real bank loans can depend on  $(y, \pi, \iota)$

| Parameter      | Description                    | Parameter Value     | SE   |
|----------------|--------------------------------|---------------------|------|
| h <sub>0</sub> | Constant term                  | -3.396              | 0.54 |
| $h_1$          | Coefficient on L               | 1.88                | 0.57 |
| $\rho_L$       | Coefficient on the lagged $L$  | .95                 |      |
| $\phi_0$       | Intercept                      | $(1 -  ho_L) * 0.2$ |      |
| $\phi_i$       | Coefficient on the policy rate | 0                   | -    |
| $\phi_y$       | Coefficient on output gap      | 0.18                | 0.04 |
| $\phi_{\pi}$   | Coefficient on inflation gap   | -1                  |      |

# Basic Trade-off



\*With  $L_0 = 0.2$  ( $L_0$  is the lagged 5-yr growth rate of real bank loans)

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# **Optimal Policy**



# Sensitivity Analysis

Crisis probability more sensitive to policy rate (higher  $h_1$  and  $\phi_y$ ).



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Optimal Policy with Parameter Uncertainty

Motivation:

 Parameters related to crises are particularly uncertain because crises are infrequent.

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Optimal Policy with Parameter Uncertainty

- 4 Sources of Uncertainty:
  - ► *h*<sub>1</sub>: Elasticity of crisis probability to credit conditions.
  - $\phi_y$ : Elasticity of credit conditions to output.
  - $(\pi_{2,c}, y_{2,c})$ : "Severity" of the crisis.
  - $(\sigma, \kappa)$ : Elasticity of today's output/inflation to policy rate.

- 2 Types of Policymaker:
  - Bayesian
  - Robust

# 2 Types of Policymakers

Bayesian policymaker:

$$\min_{i_1} E_{1,\theta} \big[ u(y_1, \pi_1) + \beta WL_2 \big]$$

Robust policymaker:

$$\min_{i_1} \left[ \max_{\theta \in [\theta_{\min}, \theta_{\max}]} u(y_1, \pi_1) + \beta E_1[WL_2] \right]$$

where  $\boldsymbol{\theta}$  is the set of parameters subject to uncertainty.

# Calibration

| Parameter            | Value            | Probability                    |
|----------------------|------------------|--------------------------------|
| Uncertain E          | lasticity of Cri | sis Prob. to Credit Conditions |
| h <sub>1,min</sub>   | 0.74             | 1/3                            |
| $h_{1,base}$         | 1.88             | 1/3                            |
| $h_{1,max}$          | 3.02             | 1/3                            |
| Uncertain S          | everity of Srisi | S                              |
| $\pi_{2,c,min}$      | -0.03/4          | 1/3                            |
| $\pi_{2,c,base}$     | -0.02/4          | 1/3                            |
| $\pi_{2,c,max}$      | -0.01/4          | 1/3                            |
| Y <sub>2,c,min</sub> | -0.15            | 1/3                            |
| Y2,c,base            | -0.1             | 1/3                            |
| У2,c,max             | -0.05            | 1/3                            |

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# Uncertain crisis prob. $(h_1)$



\**h*<sub>1</sub>: Elasticity of crisis prob. to credit conditions

#### Trade-off faced by the Bayesian Policymaker



Fig: The Effects of a Mean-Preserving Spread on  $h_1$  for the Crisis Probability Function:  $\gamma_1 = \frac{\exp(h_0 + h_1 L_1)}{1 + \exp(h_0 + h_1 L_1)}$ 



# Objective function of the hypothetical evil agent



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# Uncertain severity of the crisis $(y_{2,c} \text{ and } \pi_{2,c})$



\*( $y_{2,c}, \pi_{2,c}$ ): Output gap and inflation in the crisis

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# BayesianRobust $h_1$ HigherHigher $(\pi_{2,c}, y_{2,c})$ HigherHigher $(\sigma, \kappa)$ LowerLower

#### Table: Effects of Uncertainty on Optimal Policy Rate

#### Conclusions

- Solve for optimal interest-rate policy in a New Keynesian model with endogenous financial crises.
  - Optimal adjustment to interest rates in response to credit conditions is very small under baseline calibration.
  - Optimal policy can call for larger interest-rate adjustments under alternative/plausible calibrations.

- Compute optimal policy under parameter uncertainty.
  - Bayesian and robust-control central banks should respond more aggressively when probability and severity of financial crises are uncertain.

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Every quarter, participants in the Survey of Professional Forecasters (SPF) report the probability distribution of the growth rate of real average GDP and CPI expected over the current and next calendar years.

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Great Recession episode:

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- Great Recession episode:
  - Realized average real GDP fell by -0.29% in 2008, and -2.81% in 2009.

Every quarter, participants in the Survey of Professional Forecasters (SPF) report the probability distribution of the growth rate of real average GDP and CPI expected over the current and next calendar years.

- Great Recession episode:
  - Realized average real GDP fell by -0.29% in 2008, and -2.81% in 2009.
  - CPI inflation recorded a negative entry in 2008:Q4 and quickly reverted into positive territory.

Figure: Probability of Negative Growth of Average Real GDP in 2008 and 2009



(a) 2008

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PR(DRCPI 2008 < 0)

Figure: Probability of Negative Growth of Average CPI in 2008 and 2009



PR(DRCPI 2009 < 0)

 Supporting evidence of forecasters did not anticipate effects of the financial crisis of 2007-2009.



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 Expectations that recession will be lengthy and costly is updated with a lag to unfolding of events.



 Supporting evidence of forecasters did not anticipate effects of the financial crisis of 2007-2009.

Expectations that recession will be lengthy and costly is updated with a lag to unfolding of events.

 Forecasters' expectations for GDP growth and CPI inflation do not seem to respond preemptively to the accumulation of financial imbalances in the 2000s.



# Sensitivity Analysis (II)

Larger declines in output and inflation during a crisis (lower  $y_{2,c}$ and  $\pi_{2,c}$ ).



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# Sensitivity Analysis (III)

Today's output and inflation less sensitive to policy rate (lower  $\sigma$  and  $\kappa$ ).



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#### Calibration Crisis Probability

Predictor in annual terms:

$$L_t^a = \Sigma_{s=0}^4 \Delta \log \frac{B_{i,t-s}}{P_{i,t-s}}$$

Predictor in quarterly terms:



applied to post-war U.S. data:



# Calibration Bank Lending Growth

• Quarterly real credit growth:  $\Delta \log \frac{B_t}{P_t} = \Delta \log B_t - \pi_t$ 



#### Calibration Bank Lending Growth

- Quarterly real credit growth:  $\Delta \log \frac{B_t}{P_t} = \Delta \log B_t \pi_t$
- We estimate a process for nominal bank lending growth:

$$\Delta \log B_t = c + \phi_i i_t + \phi_y y_t + \varepsilon_t^B$$

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instrumenting  $i_t$  and  $y_t$  with their lagged values.

# Calibration Bank Lending Growth

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- We estimate a process for nominal bank lending growth:

$$\Delta \log B_t = c + \phi_i i_t + \phi_y y_t + \varepsilon_t^B$$

instrumenting  $i_t$  and  $y_t$  with their lagged values.

So that:

$$L_1 \approx \rho_L L_0 + \phi_0 + \phi_y y_1 + \pi_1 + \epsilon_1$$

| Parameter    | Description                    | Parameter Value     | SE   |
|--------------|--------------------------------|---------------------|------|
| ρ            | Coefficient on the lagged $L$  | 19/20               |      |
| $\phi_0$     | Intercept                      | $(1 -  ho_L) * 0.2$ |      |
| $\phi_i$     | Coefficient on the policy rate | 0                   | -    |
| $\phi_y$     | Coefficient on output gap      | 0.18                | 0.04 |
| $\phi_{\pi}$ | Coefficient on inflation gap   | -1                  | -    |