

Bond Risk Premia in Consumption-based Models

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Motivation

Chairman Janet Yellen: large-scale asset purchases December, 2014, Press Conference

*...we're reminding the public that we continue to hold a large stock of assets, and that is tending to push down **term premiums** in longer-term yields.*

Chairman Ben Bernanke: decomposition March, 2006, New York

*To the extent that the decline in forward rates can be traced to a decline in the **term premium**... the effect is financially stimulative and argues for greater monetary policy restraint... However, if the behavior of long-term yields reflects current or prospective economic conditions, the implications for policy may be quite different-indeed, quite the opposite.*

Chairman Alan Greenspan: conundrum June, 2005, Beijing

*That improved performance has doubtless contributed to lower inflation-related **risk premiums**, and the lowering of these **premiums** is reflected in significant declines in nominal and real long-term rates. Although this explanation contributes to an understanding of the past decade, I do not believe it explains the decline of long-term interest rates over the past year despite rising short-term rates.*

Term premium: two models & two channels

- ▶ Gaussian ATSM:
 - ▶ benchmark model
 - ▶ time-varying term premia via price of risk
 - ▶ lack micro foundation
- ▶ structural models with recursive preferences
 - ▶ Gaussian: constant term premia
 - ▶ SV: time-varying term premia via **SV**
 - ▶ economic structure imposes restrictions

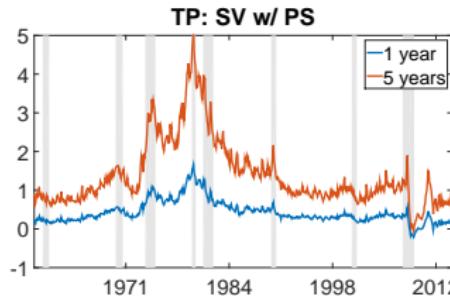
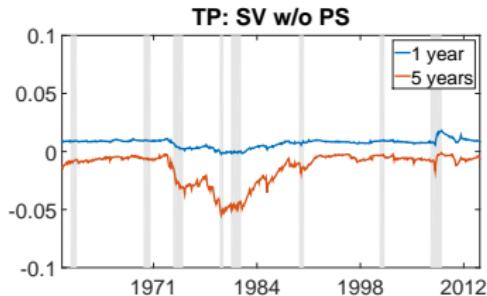
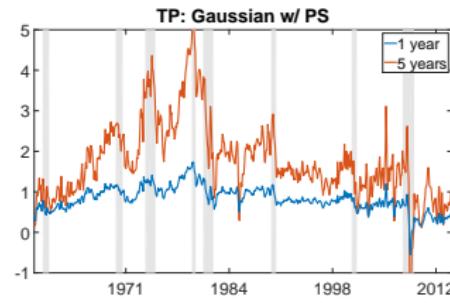
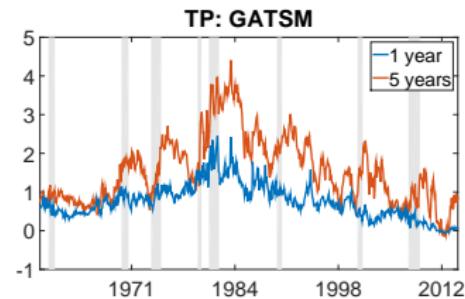
Goal of this paper: reconcile the two literatures

Contributions

- ▶ Introduce a new structural model with recursive preferences
 - ▶ preference shocks → time-varying prices of risk
- ▶ Bridge two literatures on bonds
 - ▶ GATSM and consumption based models.
- ▶ Representative agent's problem must be well-posed
 - ▶ provide conditions guaranteeing a model solution
 - ▶ representative agent cannot be too patient
 - ▶ risk aversion cannot be too large
 - ▶ the geometry of the feasible regions makes estimation challenging

▶ Literature

Results: term premia



Outline

- 1 Model
- 2 Estimation
- 3 Results
- 4 Model solution

Gaussian Model

Agent's problem

$$\begin{aligned} V_t &= \max \left[(1 - \beta) \Upsilon_t C_t^{1-\eta} + \beta \left\{ E_t \left[V_{t+1}^{1-\gamma} \right] \right\}^{\frac{1-\eta}{1-\gamma}} \right]^{\frac{1}{1-\eta}} \\ \text{s.t. } W_{t+1} &= (W_t - C_t) R_{c,t+1} \end{aligned}$$

Log stochastic discount factor

$$\begin{aligned} m_{t+1} &= \vartheta \ln(\beta) + \vartheta \Delta v_{t+1} - \eta \vartheta \Delta c_{t+1} + (\vartheta - 1) r_{c,t+1}, \\ m_{t+1}^{\$} &= m_{t+1} - \pi_{t+1} \end{aligned}$$

where $\vartheta = \frac{1-\gamma}{1-\eta}$

State variables

$$\begin{aligned} \Delta c_t &= Z'_c g_t, \quad \pi_t = Z'_\pi g_t, \\ g_{t+1} &= \mu_g + \Phi_g g_t + \Sigma_{0,g} \varepsilon_{g,t+1} \quad \varepsilon_{g,t+1} \sim N(0, I). \end{aligned}$$

- It's a companion form, nesting long-run risk & VARMA.

Preference shocks

$$\begin{aligned}\Delta v_{t+1} &= Z_v' g_{t+1} + \Lambda_1(g_t) + \Lambda_2(g_t)' \varepsilon_{g,t+1} \\ \Lambda_2(g_t) &= -\eta \Sigma_{0,g}^{-1} (\lambda_0 + \lambda_g g_t)\end{aligned}$$

- ▶ $Z_v' g_{t+1}$: Albuquerque, Eichenbaum, & Rebelo (2014)
 - ▶ no time-varying price of risk.
- ▶ $\lambda_g \neq 0 \Rightarrow$ **time-varying price of risk/ term premium**

Sources of risk premia

$$m_{t+1}^{\$} - E_t [m_{t+1}^{\$}] = -\lambda_{g,t}^{\\$,/} \varepsilon_{g,t+1}$$

where the price of risk $\lambda_{g,t}^{\$}$ is

$$\begin{aligned} \Sigma_{0,g} \lambda_{g,t}^{\$} &= \Sigma_{0,g} \Sigma'_{0,g} (\gamma Z_c + Z_\pi) - \kappa_1 \frac{(\eta - \gamma)}{(1 - \eta)} \Sigma_{0,g} \Sigma'_{0,g} D_g \\ &\quad - \vartheta \Sigma_{0,g} \Sigma'_{0,g} Z_v + \eta \vartheta \lambda_0 + \eta \vartheta \lambda_g g_t \end{aligned}$$

recursive preferences
power utility
preference shocks

- ▶ the price of risk only varies through time if $\lambda_g \neq 0$.
- ▶ Campbell & Cochrane (1999): Risk sensitivity function $\lambda(s_t)$ creates time-varying price of risk

$$m_{t+1} - E_t [m_{t+1}] = -\gamma \sigma_c (1 + \lambda(s_t)) \varepsilon_{c,t+1}$$

- ▶ Our model

- ▶ $\Lambda_2(g_t)$ is (potentially) a function of any element of the state vector g_t .
- ▶ Bond prices are known in closed-form.

Nominal risk neutral measure $\mathbb{Q}^{\$}$

$$g_{t+1} = \mu_g^{\mathbb{Q},\$} + \Phi_g^{\mathbb{Q},\$} g_t + \Sigma_{0,g} \varepsilon_{g,t+1}^{\mathbb{Q},\$}, \quad \varepsilon_{g,t+1}^{\mathbb{Q},\$} \sim N(0, I).$$

where

$$\Sigma_{0,g} \lambda_{g,t}^{\$} = (\mu_g - \mu_g^{\mathbb{Q},\$}) + (\Phi_g - \Phi_g^{\mathbb{Q},\$}) g_t$$

Therefore,

$$\Phi_g^{\mathbb{Q},\$} = \Phi_g - \eta \vartheta \lambda_g^{\text{preference shocks}}$$

- ▶ Same form as a Gaussian ATSM.
- ▶ the price of risk only varies through time if

$$\lambda_g \neq 0 \Rightarrow \Phi_g \neq \Phi_g^{\mathbb{Q},\$}$$

- ▶ Duffee (2002): important feature of the data

Bond prices

Bond prices

$$P_t^{\$, (n)} = E_t \left[\exp \left(m_{t+1}^{\$} \right) P_{t+1}^{\$, (n-1)} \right]$$

yields

$$y_t^{\$, (n)} \equiv -\frac{1}{n} \ln \left(P_t^{\$, (n)} \right) = a_n^{\$} + b_{n,g}^{\$,'} g_t$$

- ▶ Bond loadings are similar to Gaussian ATSMs.
- ▶ They are functions of (β, γ, ψ) , where $\psi = 1/\eta$

▶ Loadings

Consumption-inflation representation

$$y_t^{\$, (n)} = -\ln(\beta) + \frac{\eta}{n} \sum_{j=0}^{n-1} E_t^{\mathbb{Q}, \$} \Delta c_{t+j+1} + \frac{1}{n} \sum_{j=0}^{n-1} E_t^{\mathbb{Q}, \$} \pi_{t+j+1} - \frac{1}{n} \sum_{j=0}^{n-1} E_t^{\mathbb{Q}, \$} \Delta v_{t+j+1} + J.I.$$

SV model: dynamics

Stochastic volatility process from Creal & Wu (2015 JEconometrics)

$$\begin{aligned} g_{t+1} &= \mu_g + \Phi_g g_t + \Phi_{gh} h_t + \Sigma_{gh} \varepsilon_{h,t+1} + \Sigma_{g,t} \varepsilon_{g,t+1} & \varepsilon_{g,t+1} &\sim N(0, I) \\ \Sigma_{g,t} \Sigma'_{g,t} &= \Sigma_{0,g} \Sigma'_{0,g} + \sum_{i=1}^H \Sigma_{i,g} \Sigma'_{i,g} h_{it} \\ h_{t+1} &\sim NCG(\nu_h, \Phi_h, \Sigma_h) \\ \varepsilon_{h,t+1} &= h_{t+1} - E_t[h_{t+1}|h_t] \end{aligned}$$

Long run risk

$$\begin{aligned} \pi_{t+1} &= \bar{\pi}_t + \varepsilon_{\pi_1,t+1} & \varepsilon_{\pi_1,t+1} &\sim N(0, h_{t,\pi_1}) \\ \Delta c_{t+1} &= \bar{c}_t + \varepsilon_{c_1,t+1} & \varepsilon_{c_1,t+1} &\sim N(0, h_{t,c_1}) \\ \bar{\pi}_{t+1} &= \mu_\pi + \phi_\pi \bar{\pi}_t + \phi_{\pi,c} \bar{c}_t + \varepsilon_{\pi_2,t+1} & \varepsilon_{\pi_2,t+1} &\sim N(0, h_{t,\pi_2}) \\ \bar{c}_{t+1} &= \mu_c + \phi_{c,\pi} \bar{\pi}_t + \phi_c \bar{c}_t + \sigma_{c,\pi} \varepsilon_{\pi_2,t+1} + \varepsilon_{c_2,t+1} & \varepsilon_{c_2,t+1} &\sim N(0, h_{t,c_2}) \end{aligned}$$

SV model: preference shock

Preference shock

$$\begin{aligned}\Delta v_{t+1} &= Z'_v g_{t+1} + \Lambda_1(g_t, h_t) + \Lambda_2(g_t, h_t)' \varepsilon_{g,t+1} + \Lambda_3(h_t)' \varepsilon_{h,t+1} \\ \Lambda_2(g_t, h_t) &= -\eta \Sigma_{g,t}^{-1} (\lambda_0 + \lambda_g g_t + \lambda_{gh} h_t), \quad \Lambda_3(h_t) = -\lambda_h.\end{aligned}$$

pricing kernel

$$m_{t+1}^{\$} - E_t[m_{t+1}^{\$}] = -\lambda_{g,t}' \varepsilon_{g,t+1} - \lambda_{h,t}' \tilde{\varepsilon}_{h,t+1}$$

risk premium

$$\begin{aligned}\Sigma_{g,t} \lambda_{g,t}^{\$} &= \Sigma_{0,g} \Sigma'_{0,g} (\gamma Z_c + Z_\pi) - \kappa_1 \frac{(\eta - \gamma)}{(1 - \eta)} \Sigma_{0,g} \Sigma'_{0,g} D_g - \vartheta \Sigma_{0,g} \Sigma'_{0,g} Z_v + \eta \vartheta \lambda_0 + \eta \vartheta \lambda_g g_t \\ &\quad + ([\gamma Z_c + Z_\pi] \otimes I_G)' \tilde{S}_g h_t - \kappa_1 \frac{(\eta - \gamma)}{(1 - \eta)} (D_g \otimes I_G)' \tilde{S}_g h_t \\ &\quad - \vartheta (Z_v \otimes I_G)' \tilde{S}_g h_t + \eta \vartheta \lambda_{gh} h_t + \text{leverage effects} \\ \Sigma'_{h,t} \lambda_{h,t}^{\$} &= \Sigma'_{gh} (\gamma Z_c + Z_\pi) - \kappa_1 \frac{(\eta - \gamma)}{(1 - \eta)} (\Sigma'_{gh} D_g + D_h) - \vartheta (\Sigma'_{gh} Z_v - \lambda_h)\end{aligned}$$

Data

Monthly data from Feb. 1959 to June 2014

Yields

- ▶ Fama-Bliss zero-coupon yields from CRSP
- ▶ maturities: 3m, 1y, 2y, 3y, 4y, 5y

Inflation + Population

- ▶ FRED database at St. Louis FRB
- ▶ CPI inflation
- ▶ Civilian population over 16

Consumption

- ▶ U.S. Bureau of Economic Analysis
- ▶ non-durables + services

Estimation approach

Step 1: estimate the time series dynamics $(\theta^{\mathbb{P}}, g_t, h_t)$

- ▶ use MCMC + particle filters → Particle Gibbs sampler. [Details](#)

Step 2: estimate $\theta^u = (\beta, \gamma, \psi)$ and θ^λ

- ▶ Run a cross-sectional regression on filtered estimates \hat{g}_t and \hat{h}_t

$$y_t^\$ = A^\$ \left(\theta^u, \theta^\lambda \right) + B_g^\$ \left(\theta^u, \theta^\lambda \right) \hat{g}_t + B_h^\$ \left(\theta^u, \theta^\lambda \right) \hat{h}_t + \eta_t \quad \eta_t \sim N(0, \Omega)$$

to estimate $\hat{\zeta}^r = \left(\hat{A}^{\\$,r}, \text{vec} \left(\hat{B}_g^{\\$,r} \right)', \text{vec} \left(\hat{B}_h^{\\$,r} \right)' \right)'$

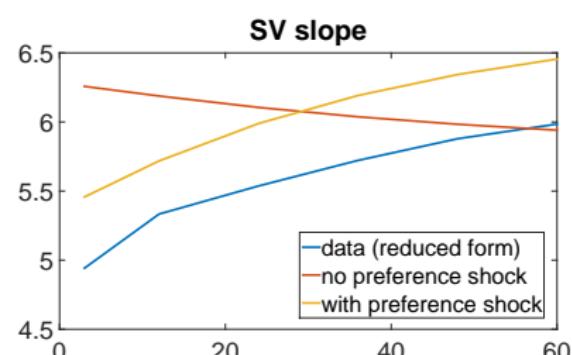
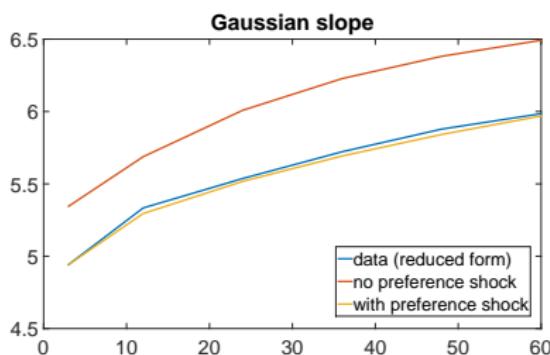
- ▶ We use minimum- χ^2 estimation of Hamilton & Wu (2012 JEconometrics)

$$\underset{\theta^u, \theta^\lambda}{\operatorname{argmin}} \quad T \left(\hat{\zeta}^r - \zeta \left(\theta^u, \theta^\lambda \right) \right)' W \left(\hat{\zeta}^r - \zeta \left(\theta^u, \theta^\lambda \right) \right)$$

Structural parameters

	Gaussian	Gaussian+PS	SV	SV+PS			
β	1.00	1.00	1.00	1.00			
ψ	0.59	1.09	0.61	0.97			
γ	84.4	416.6	1e-5	1.70			
$\bar{p}c$	7.78	5.54	7.06	8.16			
Φ_g	0.95 -0.02	0.03 0.87	0.95 -0.02	0.03 0.87	0.96 0.00	0.05 0.91	
Φ_g^Q			0.97 0.02	0.93 0.36		0.98 -0.00	-0.02 0.99
MCS	1330	477	3794	1027			

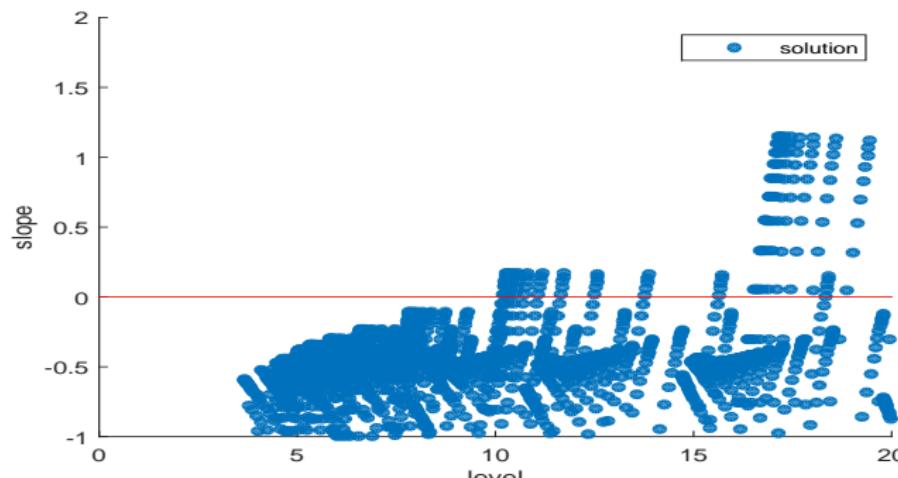
Unconditional yield curves



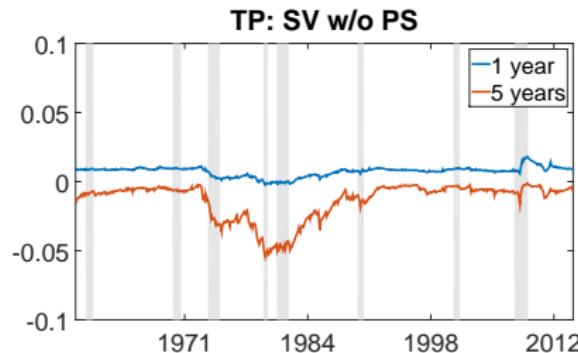
Why SV is doing worse?

- ▶ SV seems to be more flexible with h_t
- ▶ But there are more moments to match
- ▶ There are only 3 free parameters to match all
- ▶ It's difficult to match both the average level, and slope

$$\psi = [0, 20]; \beta = [0.6, 1]; \gamma = [0.001, 1100]$$



Term premia: SV without preference shock



parameters from Bansal & Yaron (2004)



Comparison with the literature

Tension between fitting macroeconomic variables and yields

- ▶ Macro variables: \mathbb{P} parameters
- ▶ Cross section of yields: \mathbb{Q} parameters
- ▶ Term premia: the difference between \mathbb{P} and \mathbb{Q}

In consumption-based models with recursive preferences, $\Phi_g = \Phi_g^{\mathbb{Q}}$

- ▶ If we force the model to fit macro variables (*ours*), then
 - ▶ Φ_g determines the slope is downward
- ▶ If we force the model to fit the upward slope (*literature*), then
 - ▶ $\Phi_g^{\mathbb{Q}}$ determines factor dynamics
 - ▶ macro factors mimic level, slope and curvature of yields

Autocovariance of volatility, $\Phi_h \neq \Phi_h^{\mathbb{Q}}$, is not enough to break the tension

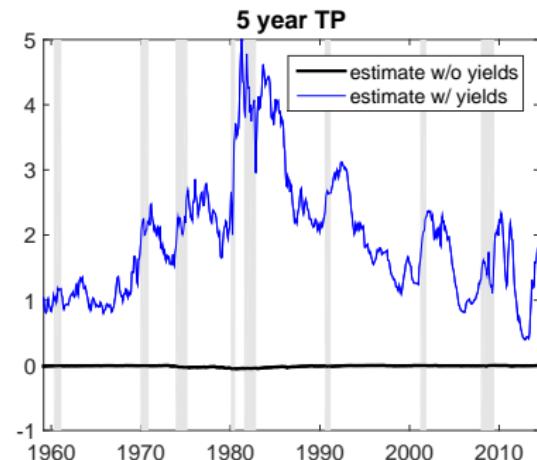
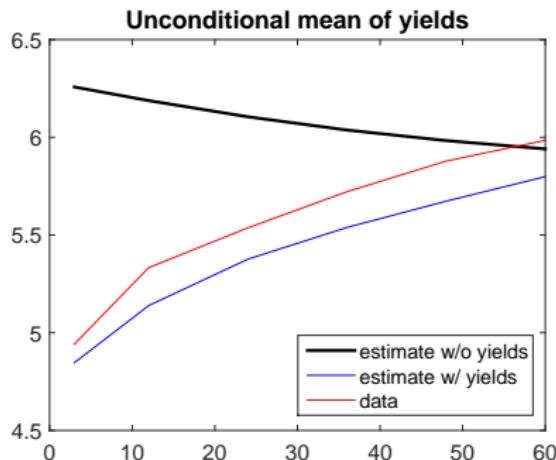
Preference shock allows $\Phi_g \neq \Phi_g^{\mathbb{Q}}$

Macro factors and yields

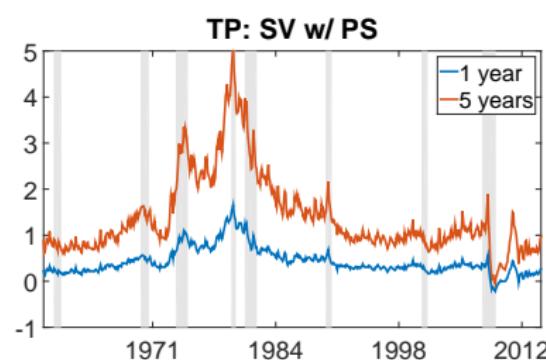
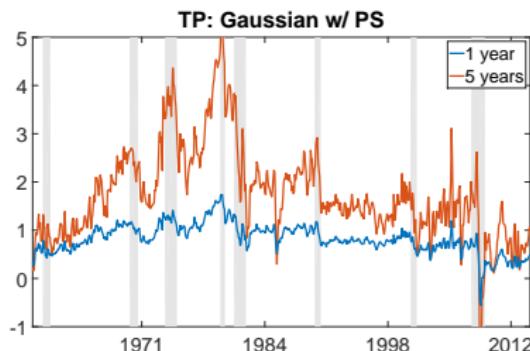
Regression R^2 s of macro factors on yields

	estimates w yields		estimates w/o yields
	w/o m.e.	w/ m.e.	
inflation	1	0.90	0.28
consumption	1	0.99	0.58
inflation V	1	0.66	0.47
consumption V	1	0.87	0.23

Slope and term premium in SV model



Adding preference shocks



- ▶ Preference shock breaks the strong tie between \mathbb{P} and \mathbb{Q}
- ▶ With the preference shock, SV does not add more flexibility for term premia.

Model solution

- ▶ Step #1: log-linearize $r_{c,t+1}$ via Campbell & Shiller (1989)

$$r_{c,t+1} = \kappa_0 + \kappa_1 pc_{t+1} - pc_t + \Delta c_{t+1}$$

- ▶ Step #2: guess a solution

$$pc_t = D_0 + D'_g g_t + D'_h h_t$$

- ▶ Step #3: solve the fixed point problem.

$$\bar{pc} = D_0(\bar{pc}) + D_g(\bar{pc})' \bar{\mu}_g + D_h(\bar{pc})' \bar{\mu}_h$$

- ▶ Step #4: plug the solution into the SDF.
- ▶ Problem: a solution to the fixed point problem does not always exist.

Solution to the fixed point problem

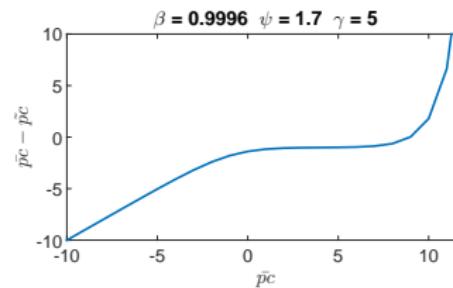
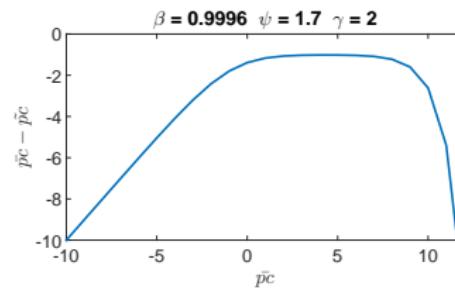
Proposition

There is a value $\bar{\beta}(\psi, \gamma, \theta^r)$ such that if $\beta < \bar{\beta}$, then there exists a solution for the fixed point problem.

Sketch of proof

Define $\tilde{p}_c (\bar{p}_c) = D_0 (\bar{p}_c) + D_g (\bar{p}_c)' \bar{\mu}_g + D_h (\bar{p}_c)' \bar{\mu}_h$

The fixed point problem has a solution if $\bar{p}_c - \tilde{p}_c (\bar{p}_c) = 0$



Solution to the fixed point problem: Gaussian

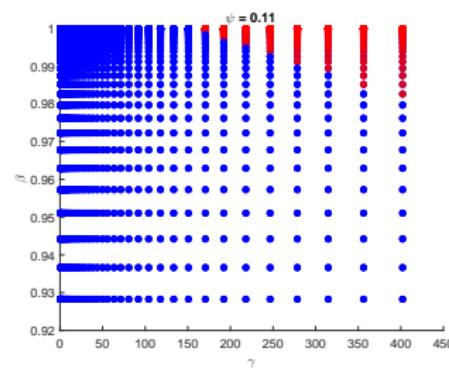
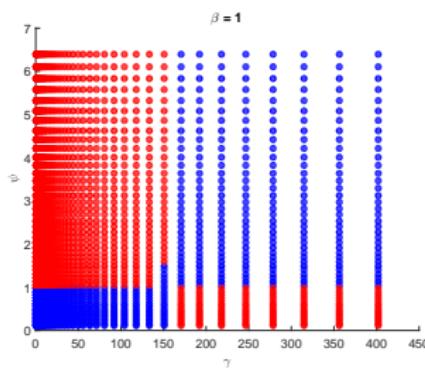
Corollary

If there is no preference shock, then

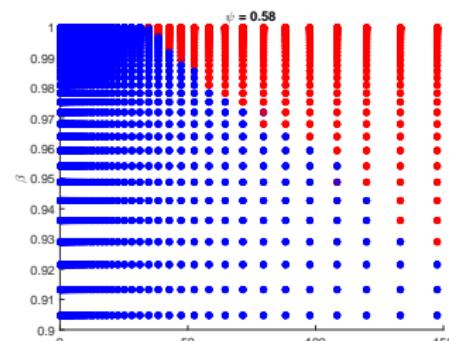
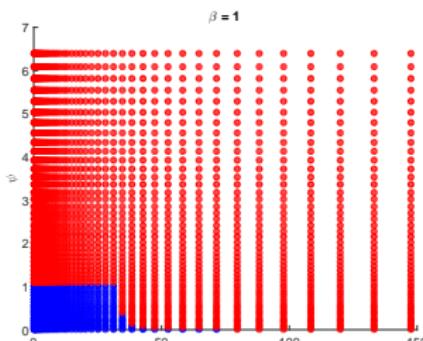
1. If $Z_1^{\infty'} \mu_g \leq 0$, then $\vartheta = \frac{1-\gamma}{1-\eta} < 0$ guarantees the existence of a solution.
2. If $\beta = 1$, there is a value $\bar{\gamma}(\theta^r)$ such that $\frac{\bar{\gamma}-\gamma}{1-\eta} < 0$ guarantees a solution.
3. For any ψ , $\bar{\beta}$ is monotonic in γ : for $\psi > 1$, then $\frac{d\bar{\beta}}{d\gamma} > 0$; for $\psi < 1$, then $\frac{d\bar{\beta}}{d\gamma} < 0$.

Model #1: long-run risk, no SV, no preference shock

Gaussian model



SV model



Conclusion

- ▶ Build a new consumption-based asset pricing model
 - ▶ capture realistic dynamics for risk premia
 - ▶ **preference shocks** → flexible time-varying prices of risk
 - ▶ stochastic volatility → time-varying quantity of risk
- ▶ Empirical results:
 - ▶ Flexible time-varying price of risk is key
 - ▶ term premia fluctuate with expected inflation
 - ▶ Term premia from recursive preferences with stochastic volatility are implausible
- ▶ Provide conditions guaranteeing a solution for these models.

Literature:

Consumption-based models

- ▶ recursive preferences: *Piazzesi Schneider (07), Le Singleton (10)*
- ▶ recursive preferences + SV: *Bansal Yaron (04), Bansal Gallant Tauchen (07), Bansal Shaliastovich (13)*,
- ▶ habit formation: *Wachter (06)*
- ▶ recursive preferences + SV in ICAPM: *Campbell Giglio et al. (14)*
- ▶ recursive preferences + preference shocks: *Albuquerque Eichenbaum Rebelo (14) Schorfheide Song Yaron (14)*

DSGE models

- ▶ habit formation: *Rudebusch Swanson (08)*
- ▶ recursive preferences: *Rudebusch Swanson (08), van Binsbergen et al. (12), Dew-Becker (14)*
- ▶ solution methods: *Caldara et al. (12)*

Literature:

Term premium

- ATSM: *Duffee (02), Ang Piazzesi (03), Wright (11), Bauer Rudebusch Wu (12)*

Model solution

- *Hansen Scheinkman (12), Campbell Giglio et al. (14)*

▶ Back

Gaussian bond loadings

$$a_n = -\frac{1}{n} \bar{a}_n, b_{n,g} = -\frac{1}{n} \bar{b}_{n,g}$$

$$\begin{aligned}\bar{a}_n^{\$} &= \bar{a}_{n-1}^{\$} + \bar{a}_1^{\$} + (\mu_g - \eta\vartheta\lambda_0 + \Sigma_{0,g}\Sigma'_{0,g}[Z_2 - Z_\pi])' \bar{b}_{n-1,g}^{\$} \\ &\quad + \frac{1}{2} \bar{b}_{n-1,g}^{\$,'} \Sigma_{0,g} \Sigma'_{0,g} \bar{b}_{n-1,g}^{\$} \\ \bar{b}_{n,g}^{\$} &= (\Phi_g - \eta\vartheta\lambda_g)' \bar{b}_{n-1,g}^{\$} + \bar{b}_{1,g}^{\$}\end{aligned}$$

where the initial conditions are

$$\begin{aligned}\bar{a}_1^{\$} &= \ln(\beta) + \bar{\Lambda} + (Z_v - \eta Z_c - Z_\pi)' (\mu_g - \eta\vartheta\lambda_0) - \frac{(\vartheta - 1)\vartheta}{2} Z_1' \Sigma_{0,g} \Sigma'_{0,g} Z_1 \\ &\quad + \frac{1}{2} Z_2' \Sigma_{0,g} \Sigma'_{0,g} Z_2 + \frac{1}{2} Z_\pi' \Sigma_{0,g} \Sigma'_{0,g} Z_\pi - Z_2' \Sigma_{0,g} \Sigma'_{0,g} Z_\pi \\ \bar{b}_{1,g}^{\$} &= (\Phi_g - \eta\vartheta\lambda_g)' (Z_v - \eta Z_c - Z_\pi)\end{aligned}$$

where

$$\begin{aligned}Z_1 &= Z_v + (1 - \eta) Z_c + \kappa_1 D_g \\ Z_2 &= \vartheta Z_v - \gamma Z_c + (\vartheta - 1) \kappa_1 D_g\end{aligned}$$

Particle Gibbs sampler

For $j = 1, \dots, M$

$$\begin{aligned}(g_{1:T}, h_{0:T})^{(j)} &\sim p\left(g_{1:T}, h_{0:T} | Y_{1:T}, \theta^{\mathbb{P},(j-1)}\right) \\ \theta^{\mathbb{P},(j)} &\sim p\left(\theta^{\mathbb{P}} | Y_{1:T}, g_{1:T}^{(j)}, h_{0:T}^{(j)}\right)\end{aligned}$$

- ▶ Draw the state variables using the particle filter, see Andrieu, Doucet, Holenstein (10).
- ▶ Use independence Metropolis-Hastings to draw the parameters $\theta^{\mathbb{P}}$.

Particle Gibbs sampler

For $t = 1, \dots, T$, run:

- ▶ For $j = 2, \dots, J$, draw from a proposal: $(h_t)^{(j)} \sim q(h_t | h_{t-1}^{(j)}, y_t, \theta)$.
- ▶ For $j = 1, \dots, J$, calculate the importance weight:

$$w_t^{(j)} \propto \frac{p(y_t | h_t^{(j)}, \theta) p(h_t^{(j)} | h_{t-1}^{(j)}, \theta)}{q(h_t^{(j)} | h_{t-1}^{(j)}, y_t, \theta)}$$

- ▶ For $j = 1, \dots, J$, normalize the weights: $\hat{w}_t^{(j)} = \frac{w_t^{(j)}}{\sum_{j=1}^J w_t^{(j)}}$.
- ▶ Conditionally resample the particles $\{h_t^{(j)}\}_{j=1}^J$ with probabilities $\{\hat{w}_t^{(j)}\}_{j=1}^J$. In this step, the first particle $h_t^{(1)}$ always gets resampled and may be randomly duplicated.

Key point: the particle approximation: $\{\hat{w}_{1:T}, h_{0:T}\}_1^J \approx p(h_{0:T} | y_{1:T}, \theta)$

▶ Back